Lecture 4 : Random variable and expectation

Study Objectives: to learn the concept of

- 1. Random variable (rv), including discrete rv and continuous rv; and the distribution functions (pmf, pdf and cdf).
- 2. Conditional expectation and conditional distribution; independence.
- 3. Expectation, variance and covariance.
- 4. Weak law of large numbers.
- 5. Moment generating function.

Random Variable: examples

- Example 1:
 X= the outcome of a fair dice ->
 We call X a discrete 'random variable'.
- -- Pr(X=1)=1/6,, Pr(X=6)=1/6;
- -- $\sum_{\{x\}} \Pr(X = x) = 1, \{x\} = \{1, 2, 3, 4, 5, 6\}$
- Example 1 (cont.) See textbook page 89~90 for a joint distribution of X=X1+X2, X1 and X2 are the outcomes of two independent fair dices.

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• Example: Two electronic components with four possible results: Pr(d,d)=0.09, Pr(d,a)=0.21, Pr(a,d)=0.21, Pr(a,a)=0.49; where d and a denote *defective* and *acceptable* respectively.



(1) If X=number of acceptable components $\Rightarrow Pr(X=0)=0.09, Pr(X=1)=0.42, Pr(X=2)=0.49$ (2) If we further define I = $\begin{cases} 1, & \text{if } X=10r \ 2\\ 0, & \text{if } X=0 \end{cases}$

 $\Rightarrow \Pr(I=1)=0.91$ $\Pr(I=0)=0.09$

Note:

If A denotes the event that at least one acceptable component is obtained, the random variable I is called the **indicator** random variable for the (occurrence of) event A, because I will equal 1 or 0 depending on whether A occurs.



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Types of random variables

• Discrete or categorical

If the possible values of X is a sequence.

Probability mass function:

p(a)=P(X=a); p(a) is positive for at most a countable numbers of values of a.=> see **Example 4.2a** (textbook, page 92.)

• Continuous

The possible values of X is an interval and is a subset of $(-\infty,\infty)$

Cumulative distribution function (cdf) $F(a) = \sum_{\{x: x \le a\}} p(x)$

• F(x), or $F_{X}(x)$,= $Pr(X \leq x)$, $X \sim F(\cdot)$

Property:

Read: X be distributed as F

$$P(a < \mathbf{X} \leq b) = P(\mathbf{X} \leq b) - P(\mathbf{X} \leq a)$$

= F(b)-F(a)

• Example: F(x)=0 for $x \le 0$ and $F(x)=1-\exp(-x^2)$ for $x>0=>P(X>1)=1-P(X\le 1)=1-\{1-\exp(-1)\}\$ = $\exp(-1)=0.368$

Probability distribution function (pdf)

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- Continuous random variable
- f(x) is called a pdf if
 - ∫_Sf(x)=1, where S is the 'support' of the random variable X, or the distribution f(•), indicating the 'range' where f(x)>0.
 - Let $I=[a,b] \subset S$, $=> \Pr(X \in I) = \int_{I} f(x) dx$
 - $dF(x)/dx|_{x=a} = f(a)$, and $F(a) = \int_{\{I=(-\infty,a)\}} f(x)dx$
- Example: [textbook page 95, example 4.2b]

Joint distribution of two random variables (continuous case)

 Joint cdf of X and Y F(x,y)=P{X≤x,Y≤y}

X and Y are both continuous rv with support $(-\infty,\infty)$

= $\int_{(-\infty,x)} \int_{(-\infty,y)} f_{\mathbf{X}}(x) f_{\mathbf{Y}}(y) dx dy$

- Of course, $\int_{(-\infty,\infty)} \int_{(-\infty,\infty)} f_X(x)f_Y(y)dxdy=1$
- The distribution of X can be obtained as:

$$F_{\mathbf{X}}(\mathbf{x}) = P\{X \leq \mathbf{x}\} = P\{X \leq \mathbf{x}, \mathbf{Y} < \infty\} = F(\mathbf{x}, \infty)$$

= $\int_{(-\infty,x)} \int_{(-\infty,\infty)} f_X(x) f_Y(y) dx dy$

Joint distribution of two random variables (discrete case)

• $p(x_i, y_j) = P\{X = x_i, Y = y_j\}$

•
$$P(X=x_i)=P(\bigcup_{j} \{X=x_i, Y=y_j\})$$

$$=\sum_{j} P\{X=x_i, Y=y_j\}$$

 $=\sum_{j} p(\mathbf{x}_{i}, \mathbf{y}_{j})$

• Example: [See textbook, page 97~98, examples 4.3a,b,c]

Independent random variables



- Continuous case: X and Y are said to be independent if P{X≦a,Y≦b}=P{X≦a}P{Y≦b}, or F_{X,Y}(a,b)= F_X (a) F_Y(b), for all a and b; or f_{X,Y}(a,b)= f_X (a) f_Y(b),
- **Discrete case:** X and Y are said to be independent if $p_{X,Y}(x,y)=p_X(x)p_Y(y)$
- Example: [Textbook page 102~104, examples 4.3d,e]

Conditional distributions



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• Example: [Textbook page 105~107, examples 4.3f,g,h]

Expectation

- E(X), or simply EX, and is common denoted as μ, is defined as
 EX=Σ {x}xP{X=x}, for discrete rv
 = ∫ {x} xfx(x)dx, for continuous rv
- Properties:
 - $E[g(x)] = \int_{\{x\}} g(x) f_{\mathbf{X}}(x) dx$ E(a**X**)= aE**X**, for any constant 'a'. E(**X**+**Y**)= E**X**+E**Y**
- Example: [Textbook page 112~118, examples 4.5a~h]

Variance

- Var(**X**)=E{(**X**- μ $)^{2}$ }, μ =E(**X**).
- **Property 1**: $Var(X) = E(X^2) \mu^2$

 $[=>Var(a\mathbf{X})=a^{2}Var(\mathbf{X})]$

For the case of two rv's X and Y with some joint distribution, we have:

- **Property 2**: E{E[X|Y]}=EX; and, E{E[g(X)|Y]}=E{g(X)}
- **Property 3**: Var(X)=E[Var(X|Y)]+Var[E(X|Y)]

Justification of Property 3

- 1. E{Var(X|Y)}=E{E(X²|Y)}-E{[g(Y)]²} by definition, where g(Y)=E(X|Y);
- 2. Var{E(X|Y)}=E{[g(Y)]²}-E(X²) can be easily obtained;
- Combining 1 and 2 we have E{Var(X|Y)}+Var{E(X|Y)}=E{E(X²|Y)}-E(X²), which is simply Var(X) using the **Property 2**.







Covariance and correlation



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- The covariance of two rv's X and Y is defined as: $cov(X,Y)=E\{(X-EX)(Y-EY)\}$, which can be easily re-expressed as E(XY)-(EX)(EY)
- Var(X+Y)=Var(X)+2cov(X,Y)+Var(Y) and Var(X-Y)=Var(X)-2cov(X,Y)+Var(Y)
- cov(aX,Y)=acov(X,Y) and cov(X,aY)=acov(X,Y)
- cov(X+Z,Y)=cov(X,Y)+cov(Z,Y)
- Propositions 4.7.2 and 4.7.3
- $\operatorname{corr}(X,Y) = \operatorname{cov}(X,Y) / \sqrt{\operatorname{(Var}(X))} \sqrt{\operatorname{(Var}(Y))}$

Moment generating functions

- The moment generating function (mgf) of the random variable **X** is defined as g(t)=E{exp(t**X**)}.
- For discrete case, $g(t) = \sum_{\{x\}} exp(t\mathbf{X})p(x)$
- For continuous case, $g(t) = \int \{x\} \exp(t\mathbf{X}) f(\mathbf{x}) d\mathbf{x}$
- Let g⁽ⁿ⁾(t)=dⁿg(t)/dtⁿ, for n>=1; then we have: g⁽ⁿ⁾(0)=E(Xⁿ), explains why g is called a mgf, and, g_{X+Y}(t)=g_X(t)g_Y(t)
- **Theorem**: The mgf uniquely determines the dsitribution; which means: a specific mgf corresponds to a unique distribution, and *vice versa*.

The weak law of large numbers (WLLN)

• Let X_1, \ldots, X_n be a set of iid rv's with mean μ . Then for any $\varepsilon > 0$,

P{ $|(X_1+...+X_n)/n-\mu| > \varepsilon$ } converges to 0.

 By Chebyshev's inequality, it can be proved that

P{ $|(X_1+...+X_n)/n-\mu| > \varepsilon$ } <Var{ $(X_1+...+X_n)/n$ }/ $\varepsilon^2 = (\sigma^2/n) \varepsilon^2$. It certainly converges to zero.

We say: (X₁+...+X_n)/n μ in probability (*in* pr); in other words, the 'sample mean' is a consistent estimator of the parameter μ, the population mean.

Homework and exercises

- 1. State and prove the Markov inequality and the Chebyshev's inequality.
- 2. Please complete the proof of Property 3:

Var(X) = E[Var(X|Y)] + Var[E(X|Y)]

3. Let $T(x_1,...,x_n)$ be unbiased for estimating μ and Var{ $T(x_1,...,x_n)$ } converges to zero when $n \uparrow \infty$. Please show that the T-statistic is consistent for μ .



4. Do the following problems in your textbook pages 130~140:
[Level 1]: 4, 6, 8, 12, 30, 36, 45, 49, 52
[Level 2]: 10, 11, 16, 28, 29, 35, 40, 50, 54, 57